

A Probabilistic Approach to a Differential-Difference Equation Arising in Analytic Number Theory

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Abstract. The differential-difference equation

$$\begin{aligned} tv'(t) + v(t - 1) &= 0, & t > 1, \\ v(t) &= 0, & t < 0, \\ v(t) &= \text{constant}, & 0 \leq t \leq 1, \end{aligned}$$

can be solved by the Monte-Carlo method, for the initial condition $v(t) = e^{-\gamma}$, $0 \leq t \leq 1$, where the $v(t)$ represent the probability density of a random variable:

$$t = \lim_{n \rightarrow \infty} \sum_{i=1}^n \prod_{j=1}^i x_j,$$

where the x_j are independent and uniformly distributed on $(0, 1)$.

I. Introduction. The function $\psi(x, y)$ is equal to the number of integers less than or equal to x and free of prime factors greater than y . Chowla and Vijayaraghavan, Ramaswami, Buchstab and de Bruijn have shown that [1]:

$$\lim_{y \rightarrow \infty} \frac{\psi(y^t, y)}{y^t} = v(t),$$

where $v(t)$ is a function satisfying

$$\begin{aligned} tv'(t) + v(t - 1) &= 0, & t > 1, \\ v(t) &= 0, & t < 0, \\ v(t) &= 1, & 0 \leq t \leq 1. \end{aligned}$$

Many authors have studied the limits and asymptotic behaviour of this equation [2]; Norton gives an exhaustive bibliography [3]. Highly accurate numerical results were obtained by Dickman, Bellman, Van de Lune ([4], [5], [6]).

The differential-difference equation solution by the Monte-Carlo method does not claim to be as accurate as these previous calculations but only shows a probabilistic aspect of this equation.

II. Stochastic Model. Let u_n be the random variable: $u_n = x_1 + x_1x_2 + \dots + x_1x_2 \dots x_n$, where x_i are independent random variables uniformly distributed on $(0, 1)$.

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It may be deduced from the distribution of a product of x_i variables that if $n \rightarrow \infty$, u_n converges in probability to a limit.

LEMMA. Assume that $v(t)$ is a function continuous on $0 < t < \infty$ satisfying the following equation:

$$(1) \quad \begin{aligned} tv'(t) + v(t - 1) &= 0, & t > 1, \\ v(t) &= 0, & t < 0, \\ v(t) &= C, & 0 \leq t \leq 1. \end{aligned}$$

This function is identical to $f(t)$: the probability density of a random variable:

$$t = \lim_{n \rightarrow \infty} \sum_{i=1}^n \prod_{j=1}^i x_j,$$

where x_i are independent random variables uniformly distributed on $(0, 1)$ if the constant C equals $e^{-\gamma}$, γ being the Euler constant.

Proof.* Introduce

$$t_a = \sum_{i=1}^{\infty} \prod_{j=1}^i x_j \quad \text{and} \quad t_b = \sum_{i=2}^{\infty} \prod_{j=2}^i x_j;$$

t_a and t_b have the same probability distribution and $t_a = x_1(1 + t_b)$, t_b and x_1 are independent.

Let $F(t)$ be the distribution function of t_a :

$$F(t) = \text{Pr}[t_a \leq t];$$

of course, if $t < 0$, then $F(t) = 0$.

If $t > 0$, we have

$$\begin{aligned} F(t) &= \text{Pr}[t_a \leq t] = \text{Pr}[x_1(t_b + 1) \leq t] \\ &= \sum \text{Pr}[t_b + 1 \leq t/x] \text{Pr}[x \leq x_1 \leq x + dx] \\ &= \sum F\left(\frac{t}{x} - 1\right) \text{Pr}[x \leq x_1 \leq x + dx] = \int_0^1 F\left(\frac{t}{x} - 1\right) dx. \end{aligned}$$

Put $(t/x) - 1 = s$, then

$$F(t) = t \int_{t-1}^{\infty} \frac{F(s)}{(s + 1)^2} ds.$$

If $0 \leq t \leq 1$, then

$$F(t) = t \int_0^{\infty} \frac{F(s) ds}{(s + 1)^2} = C \cdot t,$$

where C is a constant. Hence, $f(t) = F'(t) = C$ for $0 \leq t \leq 1$.

If $t > 1$, by differentiating once, we get

$$f(t) = (F(t) - F(t - 1))/t \geq 0;$$

by differentiating again, we find $tf'(t) = -f(t - 1)$, $t > 1$.

* I am indebted to J. J. A. M. Brands for the correction of my initial proof.

TABLE I

| n | $\mathfrak{P}r(u_n \leq 1)$ Explicit value | Monte-Carlo value (10^5 - runs) |
|----------|---|---------------------------------------|
| 2 | 0.69315 | 0.69416 |
| 3 | 0.61428 | 0.61622 |
| 4 | 0.58498 | 0.58350 |
| 5 | 0.57246 | 0.57356 |
| 6 | 0.56674 | 0.57016 |
| 7 | 0.56404 | 0.56303 |
| 8 | 0.56273 | 0.56290 |
| 9 | 0.56209 | 0.56381 |
| 10 | 0.56177 | 0.56030 |
| ∞ | 0.56146 | |

Let $h(s)$ be the Laplace transform of $f(t)$ [7]:

$$h(s) = (C_0/s) \exp\{-E_1(s)\},$$

where

$$E_1(s) = \int_s^\infty \frac{e^{-z}}{z} dz.$$

TABLE II

| t | v(t) Explicit value | t Δt = 0.1 | Monte-Carlo value (20 000 runs) | |
|-----|------------------------|---------------|---------------------------------|---|
| | | | Rough value | Smooth value using REINSCH's (10) program |
| 0 | 1 | 0 -0.1 | 0.96801 | |
| | | 0.1-0.2 | 0.99206 | |
| | | 0.2-0.3 | 1.03391 | |
| | | 0.3-0.4 | 0.96890 | |
| | | 0.4-0.5 | 1.01788 | |
| | | 0.5-0.6 | 1.01432 | |
| | | 0.6-0.7 | 1.02233 | |
| | | 0.7-0.8 | 0.99206 | |
| | | 0.8-0.9 | 1.01343 | |
| | | 0.9-1.0 | 0.95733 | |
| 1.1 | 0.9046898202 | 1.0-1.1 | 0.95911 | 0.9624 |
| 1.2 | 0.8176784432 | 1.1-1.2 | 0.91547 | 0.8874 |
| 1.2 | 0.7376357355 | 1.2-1.3 | 0.81484 | 0.8132 |
| 1.4 | 0.6635277634 | 1.3-1.4 | 0.69016 | 0.7403 |
| 1.5 | 0.5945348919 | 1.4-1.5 | 0.58419 | 0.6693 |
| 1.6 | 0.5299963708 | 1.5-1.6 | 0.57974 | 0.6006 |
| 1.7 | 0.4693717489 | 1.6-1.7 | 0.50849 | 0.5346 |
| 1.8 | 0.4122133351 | 1.7-1.8 | 0.43992 | 0.4719 |
| 1.9 | 0.3581461138 | 1.8-1.9 | 0.37492 | 0.4128 |
| | | 1.9-2.0 | 0.31169 | 0.3578 |
| 2.0 | 0.3068528194 | 2.0-2.1 | 0.29031 | 0.3070 |
| 2.1 | 0.2604057802 | 2.1-2.2 | 0.23510 | 0.2608 |
| 2.2 | 0.2203571379 | 2.2-2.3 | 0.17810 | 0.2193 |
| 2.3 | 0.1857994616 | 2.3-2.4 | 0.17098 | 0.1826 |
| 2.4 | 0.1559912639 | 2.4-2.5 | 0.16208 | 0.1506 |
| 2.5 | 0.1303195618 | 2.5-2.6 | 0.11132 | 0.1231 |
| 2.6 | 0.1082724430 | 2.6-2.7 | 0.09172 | 0.0999 |
| 2.7 | 0.08941856572 | 2.7-2.8 | 0.07748 | 0.0808 |
| 2.8 | 0.07339158076 | 2.8-2.9 | 0.05699 | 0.0653 |
| 2.9 | 0.05987811599 | 2.9-3.0 | 0.05076 | 0.0528 |
| 3.0 | 0.04860838829 | 3.0-3.1 | 0.05165 | 0.0425 |
| 3.1 | 0.03932296954 | 3.1-3.2 | 0.04186 | 0.0333 |
| 3.2 | 0.03170344451 | 3.2-3.3 | 0.02583 | 0.0250 |
| 3.3 | 0.02546472387 | 3.3-3.4 | 0.01514 | 0.0186 |
| 3.4 | 0.02037177906 | 3.4-3.5 | 0.01603 | 0.0145 |
| 3.5 | 0.01622959324 | 3.5-3.6 | 0.00980 | 0.0125 |
| 3.6 | 0.01287543418 | 3.6-3.7 | 0.01069 | 0.0121 |
| 3.7 | 0.01017283782 | 3.7-3.8 | 0.01514 | 0.0120 |
| 3.8 | 0.008006872188 | 3.8-3.9 | 0.00801 | 0.0092 |
| 3.9 | 0.006280373062 | 3.9-4.0 | 0.00534 | 0.0053 |
| 4.0 | 0.004910925648 | 4.0-4.1 | 0.00178 | 0.0018 |

Assuming that $f(t)$ is a probability $h(0) = \int_0^\infty f(t) dt = 1$, the constant C_0 equals $e^{-\gamma}$, where γ is the Euler constant.

Since $f(t) = C$ as $t = 0$, we obtain the boundary condition: $\lim_{s \rightarrow \infty} sh(s) = C = e^{-\gamma}$.

From $f(t) = C$ as $t = 1$, inverting Laplace transform, it may be deduced again that $f(1) = e^{-\gamma}$, so that

$$\begin{aligned}
 f(t) &= 0, & t < 0, \\
 f(t) &= e^{-\gamma}, & 0 \leq t \leq 1, \\
 f'(t) &= -f(t - 1)/t, & t > 1.
 \end{aligned}$$

TABLE II

| t | v(t) explicit value (*) | | $\Delta t = 0.1$ t | Monte-Carlo value (3.10 ⁵ runs) | |
|---------|-------------------------|------------------|-----------------------|--|------------------|
| 4.1 | 0.38285853 | 10 ⁻² | 4.1-4.2 | 0.39 | 10 ⁻² |
| 4.2 | 0.29754751 | 10 ⁻² | 4.2-4.3 | 0.35 | 10 ⁻² |
| 4.3 | 0.23050507 | 10 ⁻² | 4.3-4.4 | 0.27 | 10 ⁻² |
| 4.4 | 0.17799423 | 10 ⁻² | 4.4-4.5 | 0.165 | 10 ⁻² |
| 4.5 | 0.13701182 | 10 ⁻² | 4.5-4.6 | 0.135 | 10 ⁻² |
| 4.6 | 0.10514453 | 10 ⁻² | 4.6-4.7 | 0.13 | 10 ⁻² |
| 4.7 | 0.80455901 | 10 ⁻³ | 4.7-4.8 | 0.095 | 10 ⁻² |
| 4.8 | 0.61395778 | 10 ⁻³ | 4.8-4.9 | 0.065 | 10 ⁻² |
| 4.9 | 0.46728046 | 10 ⁻³ | 4.9-5.0 | 0.085 | 10 ⁻² |
| 5.0 | 0.35472534 | 10 ⁻³ | 5.0-5.1 | 0.08 | 10 ⁻² |
| (Δ) 5.1 | 0.268580 | 10 ⁻³ | | | |
| 5.2 | 0.202822 | 10 ⁻³ | | | |
| 5.3 | 0.152768 | 10 ⁻³ | | | |
| 5.4 | 0.114775 | 10 ⁻³ | | | |
| 5.5 | 0.860192 | 10 ⁻⁴ | | | |
| 5.6 | 0.643153 | 10 ⁻⁴ | | | |
| 5.7 | 0.479771 | 10 ⁻⁴ | | | |
| 5.8 | 0.357089 | 10 ⁻⁴ | | | |
| 5.9 | 0.265188 | 10 ⁻⁴ | | | |
| 6.0 | 0.196503 | 10 ⁻⁴ | | | |

(*) Calculated by 4th order TAYLOR's expansion
 (Δ) Calculated by 5th order TAYLOR's expansion

DICKMAN result Monte-Carlo value

$$\int_0^{\infty} \frac{v(t)}{(1+t)^2} dt = 0.62433 = 06238$$

III. Numerical Calculations. For $t \leq 4$, the solution of Eq. (1) is obtained by explicit expression (see Appendix); for $t > 4$, it is impossible to express the solution by means of known functions. This explicit expression can thus be used for the well-known equation of the statistic theory of damage [8].

$$\begin{aligned} tu'(t) &= u(t - 1), & t > 1, \\ u(t) &= 0, & t < 0, \\ u(t) &= 1, & 0 \leq t \leq 1. \end{aligned}$$

For $t \leq 4$, the function $v(t)$ can be calculated with an accuracy depending solely on the polylogarithms which are used in its expression [9]. The random variable u_n is very easy to simulate by means of the pseudo-random numbers of Lehmer's method.

It can be seen in Section II that the u_n distributions achieve rapid convergence as n increases.

For the calculations, n is chosen so that we cannot discriminate between the distributions of u_n and u_{n-1} because the statistical fluctuations of the pseudo-random numbers are greater than the discrepancy between them.

IV. Results. Table I gives an illustration of Section II; notice that we get the Euler constant simulated by $-\text{Log}|\Phi r[u_n \leq 1]|$, $n \rightarrow \infty$.

Table II represents the calculation of the function $v(t)$ explicitly and by simulation. Results are smoothed by the spline method [10]. Polylogarithms can be calculated by means of Chebyshev's polynomial expansion [11], [12]; Kölbig gives an excellent algorithm for the dilogarithm's calculation [13].

V. Conclusion. The main purpose of this paper is to test the ability of the Monte-Carlo method to resolve differential-difference equations, and, using a classical example, to justify further studies in the field of the statistical theory of damage and neutron transport problems [14] which involve the same mathematical data.

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Appendix.

$v(t)$ explicit behaviour,

$$v(t) = 1 - \text{Log } t, \quad 1 \leq t \leq 2,$$

$$v(t) = 1 - \text{Log } t + \left[\frac{1}{2} \text{Log}^2 t + L_2(1/t) + L_2(-1)\right], \quad 2 \leq t \leq 3,$$

$$v(t) = 1 - \text{Log } t + \left[\frac{1}{2} \text{Log}^2 t + L_2(1/t) + L_2(-1)\right]$$

$$- \left\{ \frac{1}{4} \left[L_3\left(\frac{1}{4}\right) - L_3\left(\frac{1}{(t-1)^2}\right) \right] - \frac{1}{3} (\text{Log}^3(t-1) - \text{Log}^3 2) \right.$$

$$+ \frac{1}{2} (\text{Log}^2(t-1) \text{Log } t - \text{Log}^2 2 \text{Log } 3) + L_2\left(\frac{1}{t-1}\right) \text{Log } \frac{t}{t-1}$$

$$- L_2\left(\frac{1}{2}\right) \text{Log} \left(\frac{3}{2}\right) - L_2\left(-\frac{1}{t-1}\right) \text{Log}(t-2) + L_2(-1) \text{Log } \frac{t}{3}$$

$$+ \left\{ \underbrace{\left[\text{Log } \frac{1}{2} - \text{Log } \frac{t-2}{t-1} \right]}_{V_1} + \left[\frac{1}{2} - \frac{1}{t-1} \right] \right\}$$

$$- \frac{1}{2^2} \left[\underbrace{V_1 + \frac{1}{2} \left(\frac{1}{2^2} - \frac{1}{(t-1)^2} \right)}_{V_2} \right]$$

$$+ \cdots + \frac{(-1)^{p+1}}{(p+1)^2} \left[\underbrace{V_p + \frac{1}{p+1} \left(\frac{1}{2^{p+1}} - \frac{1}{(t-1)^{p+1}} \right)}_{V_p} + \cdots \right], \quad 3 \leq t \leq 4.$$

By means of Newton's method, the explicit expression permits easy calculation of the roots t_k

$$v(t_k) = \frac{1}{k}, \quad k = 4, 5, \dots, 203.$$

For example, the roots

$$t_4 = 2.1245966, \quad t_5 = 2.2571089$$

are used by Davenport and Erdős [15].

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